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Veterinary Parasitology



A systematic review and meta-analysis of impact of strongyle parasitism on growth rates in young cattle

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ABSTRACT

Background: To identify any universal impact of strongyle parasite burden on the growth rate of young cattle. *Methods:* A meta-analysis and meta-regression of the relationship between differences in strongyle parasite burden between cohorts and average daily weight gain was conducted. Publications were identified from a search of databases applying PRISMA 2020 principles. Eligible studies had at least two groups of growing cattle on the same farm that were identical in composition, management and diet except for parasite exposure and were subject to body weight gain or average daily gain and faecal egg count measurements across the common growing period. The reference group had the lowest growth-period faecal egg count. A meta-regression estimated the impact of strongyle parasitism. The dependent variable was the log of the ratio of average daily gain between comparison groups and the reference group with the predictor variable as the common logarithm of the difference in average faecal egg count (plus 1) between the comparison and the reference groups.

Results: 26 publications containing 85 groups and 59 comparison ratios were analysed. Papers included representatives from dairy and beef industries and from pasture and feedlot production systems and from all cattleproducing continents. The comparison group average daily growth rate was 0.89 (95%CI 0.81–0.97) that of the reference group. Meta-regression identified a 0.131 linear reduction in average daily weight gain ratio for every \log_{10} increase in the difference between comparison and reference group faecal egg count. Direction of effect was consistent across all subset analyses (continent, class of stock and production system). Whilst small faecal egg count differences between the comparison and reference groups often provided similar rates of daily weight gain, the trend was negative with most comparison groups having lower daily weight gains than their reference group.

Conclusions: Strongyle parasitism of growing cattle as measured by faecal egg count is associated with reduced growth across all production systems, geographies and classes of cattle.

1. Introduction

Strongyle parasitism of growing cattle can produce disease and a reduction in productivity (Bisset, 1994; Charlier et al., 2009; Rashid et al., 2019; Sykes, 1994). Productivity impacts include impaired weight gain, reproduction, lactation, feed use efficiency and death (Corwin, 1997). Further, parasite burdens are associated with changes in behaviour of cattle including time spent walking, lying and feeding that indicate discomfort, with welfare implications, even in subclinically-affected cattle (Högberg et al., 2021).

Faecal egg counts are a rapid method of evaluating parasite burdens

in grazing animals (Nielsen, 2021). Close correlations between FEC and total worm burden in young cattle have been demonstrated in both warm and cool climates (Bryan and Kerr, 1989; Teixeira et al., 2021). However, other studies suggest that in temperate seasonal grazing systems, worm egg counts, while useful, do not accurately reflect worm burden of animals (Eysker and Ploeger, 2000).

A strong correlation has also been demonstrated between the number of larval worms taken in by cattle on pasture and their average daily gain (Burggraaf et al., 2007), but due to the cost and technical difficulty of conducting pasture larval counts, these are rarely used to make treatment or management decisions for cattle herds (Molento et al.,

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Review article



2016).

Charlier estimated the net difference in average daily gain between treated and untreated first-season grazing calves was 150–315 g/head/ day (Charlier et al., 2014). However, precise signals for determining the impact or potential economic benefit of treating or not treating young cattle have been unclear, despite the use of various measures including worm egg counts, liveweight measurements and a combination of these (Kenyon and Jackson, 2012).

Parasite treatment and control costs add to (Burggraaf et al., 2007) the economic impact of parasitism (Rashid et al., 2019). There is increasing resistance to anthelmintics (Waller, 2003), and the rate of development of new anthelmintics has slowed since the 1980 s (McKellar and Jackson, 2004), making effective control of parasites in commercial farming challenging (Sangster, 2001). Cattle farming using pastures requires effective and sustainable ways to control strongyle parasitism, especially in young stock (Sutherland and Bullen, 2015). The targeted treatment of livestock combined with maintenance of a pool of untreated (and hopefully susceptible) parasites in refugia has gained much acceptance (Charlier et al., 2009). Novel (non-chemical) controls are of increasing interest (Ketzis et al., 2006). Worm egg counts can be used to predict the level of future pasture larval contamination (Molento et al., 2016; Shaw et al., 1998). The role of grazing management in controlling host-parasite interactions is of increased importance as a management strategy (Hutchings et al., 2003).

This review examines the relationship between strongyle parasite burden (as measured by faecal egg count (FEC)) and average daily weight gain (ADG) in young, growing cattle from across different breeds, current production systems, countries, and between within-farm comparison groups using a meta-analysis. The objectives are to see if a metaregression can effectively quantify the impact of strongyle parasite burden on growth performance of farmed cattle in a generalisable way, and to estimate impact of modest worm burden on average daily weight gain under modern farming practices. This information is important in that it guides decision making on investment into parasite controls.

2. Material and methods

2.1. Eligibility criteria

The population of interest was weaned eligible cattle less than two years of age (i.e., growing) within current (from 1990 onwards) commercial farming systems, whose weight was measured alongside monitoring for parasitism with strongyles using FEC during the growth period.

2.2. Database search and systematic review

The Web of Science portal was used to search multiple reference databases (including CAB Abstracts, BIOSIS Citation Index, Current Contents, MEDLINE and SciELO Citation Index) supplemented by an identical search using Google Scholar was undertaken in September 2021. The search query was: (calf or calves or yearling* or heifer or steer or cow* or cattle) AND (FEC or "fecal egg count*" or "faecal egg count*") AND ("weight gain" or "average daily weight gain" or "average daily gain" or ADG or "body weight" or "liveweight gain" or growth or "weaning weight" or weight* or gain" or liveweight") limited to a publication year of 1990 or above.

2.3. Study selection

Publications that measured both animal weight or growth and strongyle parasite burden using FEC during the growing period in this class of cattle were suitable for inclusion in the meta-analysis. The title of publication was examined for relevance. The abstracts of publications potentially meeting eligibility criteria were then examined for compliance with the eligibility criteria. Specifically, the abstracts reporting weights and/or ADGs and strongyle parasite burden measurements of study groups identified papers for examination.

At least two groups of growing cattle were required, which was necessary for the calculation of a) the relative growth performance (ratio of ADG) of each study group to a within-study reference group; and b) the predictor variable, being the difference in FEC between the study and reference group. This approach was used to control for other studylevel (confounding) variables that may also impact growth rate — such as breed, production system, supplementary feeding and use of hormone growth promoters — thereby support pooling of results from across studies. The use of anthelmintics was not an eligibility criterion. The strongyle parasite burden was able to be estimated if there was at least one FEC taken from each group within the growing period.

The use of ADG provides partial control over differing study durations between studies under the assumption that the ratio in ADG between similarly aged cattle under near-identical conditions is centred on 1 (but can be greater or less than 1). The first weight measurement coinciding with when a FEC was taken was used where possible. If multiple FECs were recorded between weight measurements, the arithmetic average count across each measurement point within the growing period was calculated and this was used in analysis.

Papers reporting more than one study group under similar management and presenting group average weight gain along with the standard deviation of daily weight gain or a standard error of the mean of average daily weight gain and who provided group-level FECs during the growing period were forwarded for meta-analysis. Publications that presented only a pooled standard deviation/standard error from combined study group results were excluded. Studies with multiple replicates, such as farm-level replicates of the study, had each replicate uniquely identified to ensure the most appropriate reference group for ADG ratio and difference in FECs was used.

Publications were assessed for bias in allocation to groups, measurement, losses and attrition and attribution of effect (such as from differences between groups potentially due to confounding exposures, such as differential feeding rates between groups). Publications with non-random allocation to groups, excessive losses and confounding between groups were removed from analysis. Blinding of participants was not considered feasible for most field parasitological studies so this was not used as an exclusion criterion.

2.4. Statistical analysis

The ratio of means method described by Friedrich was used for the meta-regression (Friedrich et al., 2008). The natural logarithm of the ratio of ADG for each comparison group to the reference group within each study (the group with the lowest FEC for the growth period) was the dependent variable for meta-regression. The lower bound of the raw ratio is zero whereas the logarithm of the ratio can extend between negative and positive infinity and so is more suitable for linear regression. The standard deviation of the natural logarithm of the ratio of means was approximated as the sum of the standard deviations of ADG for the comparison and reference group. The key advantage of the ratio of means method is the normalisation of group performance within individual studies (scaled to a common reference ratio of 1 when there is an equivalent growth rate between the comparison and reference group). This approach helps control study-level sources of variation that influence raw growth rate (such as breed, production system or use of hormonal growth promotants) because the ratio of means captures relative performance of each group against the study reference group.

The common logarithm (base 10) difference in FEC between the study and reference group was used as the predictor variable. The difference in FEC between comparison and reference group was chosen over the ratio of FECs because the difference in counts better captures biological effect magnitude than a ratio. The difference in FEC between groups also partly controls for any measurement error arising from technique for estimating egg counts as both groups were assessed using

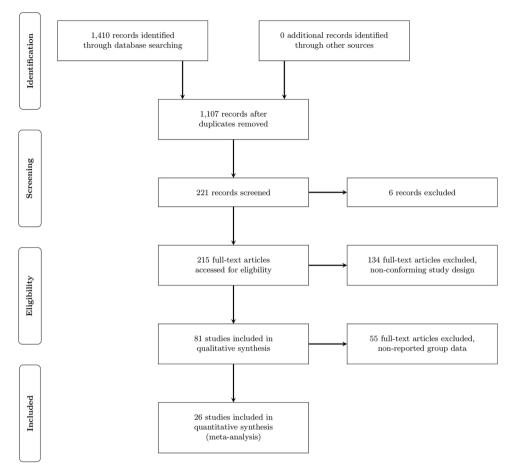


Fig. 1. PRISMA flow chart.

the same method. The common logarithm of the difference between FECs was used because this scale is more amenable to effective communication of results.

The Calc spreadsheet of LibreOffice was used for data entry and basic data manipulation (Foundation, 2020). The R Language and Environment for Statistical Computing was used for analysis (R Core Team, 2012). The R library meta was used for meta-analysis and meta-regression (Balduzzi et al., 2019). Statistical significance was set at p < 0.05.

Comparison group (within study) and study were included as random effects in meta-analysis. Heterogeneity was examined by inspection of funnel plots and examination of the O statistic. The duration of the growth period and any interaction with FEC difference were forwarded to the meta-regression model to adjust for potential confounding effects of developing natural immunity to strongyle infection with increasing age in cattle across studies of differing duration. Linear and curvilinear relationships between predictor and dependent variable were examined for continuous predictor variable. Predictor variables with a statistically significant Wald statistic for the beta coefficient were retained in the model. Predictor variables comprising part of an interaction term with a statistically significant Wald statistic for the beta coefficient were also retained in the model even if the Wald statistic for variables outside of the interaction terms was not statistically significant. A subset analysis of meta-analysis and meta-regression by continent of study was undertaken to examine geographical difference in the relationship between ADG ratio and FEC difference between groups. Linear regression was used to test for funnel plot asymmetry of the final meta-regression model.

Table 1

Publication search and filtering for meta-analysis.

| Reason | No. publications |
|--------------------------------------|------------------|
| Total records identified | 1410 |
| Total unique publications identified | 1107 |
| Title/abstract not suitable | 885 |
| Paper not attainable | 6 |
| Study design inappropriate | 134 |
| No FECs from growth period | 1 |
| No group ADG SD/SEM reported | 50 |
| Confounded comparator | 2 |
| Errors in reported results | 3 |
| Included in meta-analysis | 26 |

3. Results

The Web of Science and Google Scholar searches identified 1410 records of which there were 1107 unique publications whose title and/or abstract suggested they may be useful. Further review of the abstracts identified 221 publications for further examination. Of these, 215 were obtained, of which 81 met study design eligibility requirements. One publication did not have adequate growth-period FECs, and a further 50 did not adequately record group live weight variance/standard deviation and/or provide a standard error of the mean for each group. There was unbalanced confounding in two publications (dissimilar exposures between comparison and reference groups), and three papers likely had errors in reported estimates. A total of 26 publications providing 85 groups satisfied all inclusion criteria and were sent forward for metaanalysis. The *Preferred Reporting Items for Systematic Reviews and Meta-Analyses* (PRISMA) flow chart is presented in Fig. 1 with extra detail on

Table 2

Meta-analysis publication summary.

| Manuscript | Country | Study description | No. Comp. | Study type | | |
|-----------------------------------|--------------|----------------------------------|--------------|------------------------------|--|--|
| Kennedy (| Canada | Pasture-based | 1 | Treatment | | |
| Kennedy, 1990) | | beef cow/calf | | (anthelmintic); | | |
| | | field trial | | negative control | | |
| Boyles (Boyles | USA | Feedlot beef | 1 | Treatment | | |
| et al., 1993) | | field trial | | (anthelmintic); | | |
| Durais (Durais | 110.4 | Desture have d | - | negative control | | |
| Purvis (Purvis et al., 1994) | USA | Pasture-based beef heifer | 5 | Treatment (anthelmintic): | | |
| et al., 1994) | | field trial | | negative control | | |
| Larson (Larson | USA | Pasture-based | 1 | Treatment | | |
| et al., 1995) | | beef heifer | | (anthelmintic); | | |
| | | field trial | | negative control | | |
| Williams (Williams | USA | Pasture-based | 1 | Treatment | | |
| et al., 1995) | | cross beef | | (anthelmintic); | | |
| | | heifer field | | negative control | | |
| | D 1 · | trial | | | | |
| Agneessens (| Belgium | Pasture-based | 1 | Prospective cohor | | |
| Agneessens et al., 1997) | | beef cow/calf field trial | | (observational) | | |
| Fernandez (| USA | Feedlot beef | 2 | Treatment | | |
| Fernandez et al., | USA | steer field trial | 2 | (anthelmintic); | | |
| 1998) | | steer neid thai | | positive + | | |
| , | | | | negative controls | | |
| Schunicht (| Canada | Feedlot beef | 1 | Treatment | | |
| Schunicht et al., | | yearling field | | (anthelmintic); | | |
| 2000) | | trial | | positive control | | |
| Yamane (Yamane | Japan | Pasture-based | 1 | Treatment | | |
| et al., 2000) | | dairy heifer | | (anthelmintic); | | |
| | | field trial | | positive control | | |
| Eleonor (Eleonor | Canada | Pasture-based | 2 | (acaricide) | | |
| Elsener (Elsener et al., 2001) | Canada | dairy heifer | 2 | Treatment (anthelmintic); | | |
| et al., 2001) | | field trial | | positive control | | |
| Epperson (| USA | Pasture-based | 1 | Treatment | | |
| Epperson et al., | | beef heifer | - | (anthelmintic); | | |
| 2001) | | field trial | | negative control | | |
| Cleale (Cleale | USA | Pasture-based | 4 | Treatment | | |
| et al., 2004) | | beef heifer and | | (anthelmintic); | | |
| | | steer field trial | | negative control | | |
| Waruiru (Waruiru, | Kenya | Pasture-based | 1 | Treatment | | |
| 2004) | | cross dairy | | (anthelmintic); | | |
| | | heifer field trial | | positive control | | |
| Mertz (Mertz et al., | USA | Pasture-based | 11 | (nutrition) Treatment | | |
| 2005) | 05/1 | beef field trial | 11 | (anthelmintic); | | |
| 2000) | | beer neid that | | negative control | | |
| Reinhardt (| USA | Feedlot beef | 2 | Treatment | | |
| Reinhardt et al., | | heifer field | | (anthelmintic); | | |
| 2006) | | trial | | positive control | | |
| O'Shaughnessy (| Ireland | Pasture-based | 1 | Treatment | | |
| O'Shaughnessy | | beef cow/calf | | (anthelmintic); | | |
| et al., 2014) | | field trial | | positive control | | |
| Fazzio (Fazzio | Argentina | Feedlot beef | 2 | Treatment | | |
| et al., 2016) | | calf trial | | (anthelmintic); | | |
| Moulin (Moulin | From a a | Desture beesd | 2 | positive control | | |
| Merlin (Merlin et al., 2016) | France | Pasture-based dairy heifer | 3 | Cohort (observational) | | |
| et al., 2010) | | field trial | | (Observational) | | |
| Edmonds (| USA | Pasture-based | 3 | Treatment | | |
| Edmonds et al., | | beef steer field | | (anthelmintic); | | |
| 2018) | | trial | | positive + | | |
| | | | | negative controls | | |
| Hoglund (Hoglund | Sweden | Pasture-based | 2 | Cohort (exposure) | | |
| et al., 2018) | | dairy and | | | | |
| | | dairy-beef | | | | |
| | 0 1 | cross field trial | 1 | O-h-++(| | |
| T - 1 (T - 1 | Sweden | Pasture-based mixed dairy | 1 | Cohort (exposure) | | |
| | | mixed dairy | | | | |
| Hogberg (Hogberg et al., 2019) | | | | | | |
| Hogberg (Hogberg et al., 2019) | | and dairy-beef | | | | |
| et al., 2019) | Argentina | and dairy-beef cross field trial | 4 | Treatment | | |
| | Argentina | and dairy-beef | 4 | Treatment (anthelmintic); | | |

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Table 2 (continued)

| Manuscript | Country | Study description | No. Comp. | Study type |
|---|---------|--|--------------|--|
| das Neves (das Neves et al., 2020) | Brazil | Pasture-based dairy-beef cross cow/calf field trial | 2 | positive + negative controls Treatment (anthelmintic); positive + negative controls |
| Dudley (Dudley and Smith, 2020) | USA | Pasture-based dairy heifer field trial | 1 | Treatment (anthelmintic); negative control |
| Kasimanickam (Kasimanickam and Kasimanickam, 2021) | USA | Pasture-based beef field trial | 4 | Treatment (anthelmintic); positive control |
| Hogberg b ((Hogberg et al., 2021)) | Sweden | Pasture-based dairy steer field trial | 1 | Cohort (exposure) |

Table 3

Meta-analysis publication bias assessment.

| Manuscript | Allocation | Losses (attribution bias) |
|---------------|--|---|
| Kennedy | Not described | Nil |
| Boyles | Random assignment | Nil |
| Purvis | Random assignment | Nil |
| Larson | Stratified allocation | Nil |
| Williams | Randomised block allocation | Nil |
| Agneessens | Random assignment to exposure group | Nil |
| Fernandez | Randomised block | Nil |
| Schunicht | Random assignment | Reported, but no significant difference in losses between groups |
| Yamane | Randomised block allocation | Nil |
| Elsener | Randomised block allocation | Nine accidental losses from negative control group |
| Epperson | Randomised stratified allocation | 3 animals lost ear tags each per group - excluded from analysis |
| Cleale | Randomised block allocation | 2 animals from separate sites lost from treatment group |
| Waruiru | Weight based (block) allocation | Not reported |
| Mertz | Random assignment | Not reported |
| Reinhardt | Random assignment | Reported, but no difference between groups in mortality rate |
| O'Shaughnessy | Random assignment | Removals reported but groups not stated |
| Fazzio | Random assignment | Not reported |
| Merlin | Cohorts determined by exposure within farm | Not reported |
| Edmonds | Randomised block allocation | Not reported |
| Hoglund | Block allocation | Not reported |
| Hogberg | Block allocation | Not reported |
| Canton | Randomised block allocation | Not reported |
| das Neves | Block allocation | Not reported |
| Dudley | Randomised block allocation | Not reported |
| Kasmanickam | Randomised block allocation | Not reported |
| Hogberg b | Block allocation | Not reported |

paper exclusions in Table 1. The studies included in the meta-analysis are summarised in Table 2 and the bias assessment presented in Table 3. Given most parasitological studies require management of co-horts at group level, blinding of participants was not considered feasible for most studies so this potential bias was not examined.

There were many potentially useful publications that could not be

| Author | g | SE | Risk Ratio | CI |
|--|--------|------------------|--|---|
| Agneessens et al. 1997 | -0.07 | 0.2846 | | 0.93 [0.53; 1.62] |
| Boyles et al. 1993 | -0.05 | 0.7603 | | 0.95 [0.21; 4.23] |
| Canton et al. 2020 | | 0.2423 | | 0.87 [0.54; 1.39] |
| Canton et al. 2020 | | 0.2604 | | 0.82 [0.49; 1.36] |
| Canton et al. 2020 | | 0.2296 | | 0.55 [0.35; 0.86] |
| Canton et al. 2020 | | 0.1390 | | 0.48 [0.37; 0.64] |
| Cleale et al. 2004-G1 Cleale et al. 2004-G2 | | 0.4949 0.4949 | | 0.98 [0.37; 2.60] 0.80 [0.30; 2.10] |
| Cleale et al. 2004-G2 | | 0.4949 | | 0.78 [0.39; 1.56] |
| Cleale et al. 2004-G4 | | 0.2828 | | 0.65 [0.37; 1.13] |
| das Neves et al. 2020 | | 0.5098 | | 0.97 [0.36; 2.63] |
| das Neves et al. 2020 | | 0.5942 | ł | 1.01 [0.32; 3.25] |
| Dudley et al. 2020 | -0.30 | 0.3073 | | 0.74 [0.41; 1.35] |
| Edmonds et al. 2018 | | 0.9287 | | 0.83 [0.13; 5.14] |
| Edmonds et al. 2018 | | 0.7960 | | 0.96 [0.20; 4.57] |
| Edmonds et al. 2018 | | 0.7960 | | 0.96 [0.20; 4.57] |
| Elsener et al. 2001 | | 0.4300 | | 0.48 [0.21; 1.13] |
| Elsener et al. 2001 Epperson et al. 2001 | | 0.4300 0.0991 | | 1.30 [0.56; 3.03] 1.12 [0.92; 1.36] |
| Fazzio et al. 2016 | | 0.7277 | | 0.82 [0.20; 3.40] |
| Fazzio et al. 2016 | | 0.7454 | | 0.94 [0.22; 4.07] |
| Fernandez et al. 1998 | | 0.2827 | <u> </u> | 0.96 [0.55; 1.66] |
| Fernandez et al. 1998 | | 0.2827 | <u></u> | 0.90 [0.52; 1.57] |
| Hogberg et al. 2019 | -0.16 | 0.0520 | | 0.85 [0.77; 0.94] |
| Hogberg et al. 2021 | | 0.1423 | | 1.01 [0.76; 1.33] |
| Hoglund et al. 2018-G1 | | 0.3200 | | 0.73 [0.39; 1.36] |
| Hoglund et al. 2018-G2 | | 0.4200 | | 0.61 [0.27; 1.39] |
| Kennedy 1990 | | 0.2684 | | 0.96 [0.57; 1.62] |
| Larson et al. 1995 Merlin et al. 2016-G1 | | 0.1724 0.2400 | | 0.91 [0.65; 1.28] |
| Merlin et al. 2016-G1 | | 0.2400 | | 0.90 [0.57; 1.45] 1.06 [0.70; 1.60] |
| Merlin et al. 2016-G2 | | 0.1800 | | 1.23 [0.87; 1.75] |
| Mertz et al. 2005-G1 | | 0.2332 | <u> </u> | 0.96 [0.61; 1.52] |
| Mertz et al. 2005-G10 | | 0.2856 | | 0.92 [0.53; 1.61] |
| Mertz et al. 2005-G11 | -0.04 | 0.2560 | | 0.96 [0.58; 1.59] |
| Mertz et al. 2005-G2 | | 0.2774 | | 0.96 [0.56; 1.66] |
| Mertz et al. 2005-G3 | | 0.2545 | <u> </u> | 0.96 [0.58; 1.57] |
| Mertz et al. 2005-G4 | | 0.2048 | | 0.89 [0.60; 1.33] |
| Mertz et al. 2005-G5 | | 0.2575 0.2914 | | 0.88 [0.53; 1.46] |
| Mertz et al. 2005-G6 Mertz et al. 2005-G7 | | 0.2914 | | 0.98 [0.55; 1.74] 0.91 [0.46; 1.81] |
| Mertz et al. 2005-G8 | | 0.3375 | | 1.06 [0.55; 2.05] |
| Mertz et al. 2005-G9 | | 0.2481 | | 0.95 [0.58; 1.54] |
| O'Shaughnessy et al. 2014 | | 0.0700 | | 1.02 [0.89; 1.17] |
| Purvis et al. 1994-G1 | -0.09 | 0.1506 | - <u>è</u> - | 0.91 [0.68; 1.23] |
| Purvis et al. 1994-G2 | | 0.1574 | - <u>*</u> - | 0.89 [0.65; 1.21] |
| Purvis et al. 1994-G3 | | 0.1934 | | 0.87 [0.60; 1.28] |
| Purvis et al. 1994-G4 | | 0.1796 | - <u>1</u> - | 0.95 [0.67; 1.36] |
| Purvis et al. 1994-G5 | | 0.1988 | | 0.91 [0.62; 1.35] |
| Reinhardt et al. 2006-G1 Reinhardt et al. 2006-G2 | | 0.4704 0.7776 | | 1.00 [0.40; 2.51] |
| Schunicht et al. 2000-02 | | 1.1730 | | 1.00 [0.22; 4.59] 1.04 [0.10; 10.35] |
| Waruiru et al. 2004 | | 0.3840 | | 0.76 [0.36; 1.61] |
| Williams et al. 1995 | | 0.3090 | <u> </u> | 0.98 [0.53; 1.79] |
| Yamane et al. 2000 | | 0.0550 | | 0.76 [0.68; 0.85] |
| Random effects model Prediction interval | | | | 0.89 [0.81; 0.97] [0.81; 0.97] |
| Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$ | .0185, | p = 0.49 | | |
| | | 0. | | 0 |
| | | R | educed growth Extra growth ADG response ratio | |
| | | | | |

Fig. 2. Forest plot of average daily weight gain ratio meta-analysis.

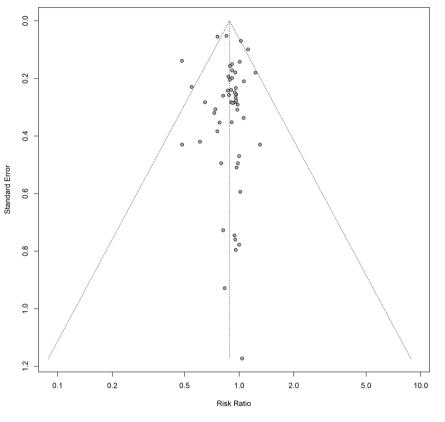


Fig. 3. Funnel plot of average daily weight gain ratio meta-analysis.

included in the meta-analysis. Losses arose primarily from failure to publish standard deviations for ADG or standard errors of the mean for ADG. Some publications did not present FECs in a form suitable for extraction and inclusion in the meta-analysis; mostly only providing FECs graphically. Not all FEC plots were sufficiently distinct to allow data to be reliably harvested using the *WebPlot Digitizer* application. Combined, this resulted in loss of studies and potentially valuable information from the meta-analysis. This was found to be an issue in other reviews of ADG (Baltzell et al., 2015).

For publications that only presented FECs graphically, the application *WebPlot Digitizer* was used to estimate the point values from plots (Rohatgi, 2021). The arithmetic average FEC for the growth period was calculated from these estimates.

Fifteen studies were undertaken in North America, with six studies conducted in Europe, three studies in South America and with Asia/ Oceania and Africa each contributing a single study providing a total of 26 studies, comprising 59 comparison ratios.

The meta-analysis forest plot is presented in Fig. 2. Groups with high strongyle burdens, as indicated by elevated FECs, had lower ADGs than groups with lower worm burdens. The mean ratio of ADG for high worm burden group in comparison to low worm burden group was 0.89 (95% CI: 0.81–0.97). There was minimal heterogeneity (Q = 53.82, on 58 degrees of freedom; p-value = 0.63). The meta-analysis funnel plot is presented in Fig. 3. The linear regression test for funnel plot asymmetry was not significant (t = 0.67, p = 0.51).

The common logarithm of the FEC difference between the comparison and reference group (plus 1) was the predictor variable of interest. The duration of the study (growth) period in months was offered to the model along with an interaction with the common logarithm of FEC difference. This was to control for any potential confounding effect of development of resistance to strongyle parasites with increasing exposure, however the duration of study was not a significant predictor. There was no significant effect of continent, industry (dairy or beef),

| Table 4 |
|---|
| Meta-regression of \log_{10} of faecal egg count difference on log of average daily |
| gain ratio |

| | Estimate | S.E. | Z-value | p-value |
|--------------------------------|----------|-------|---------|---------|
| Intercept | 0.041 | 0.063 | 0.649 | 0.52 |
| Log_{10} difference FEC (+1) | -0.131 | 0.040 | -3.310 | 0.001 |

production system (pasture or feedlot) on the outcome. The use of the logarithm of the ratio of means within study is a measure of relative performance of groups within each study and this likely controlled any effect of these between-study confounders in analysis. The final model included only the common logarithm of FEC difference (plus 1) and is summarised in Table 4. The common log of the difference in FEC between comparison and reference group (plus 1) was associated with a 0.131 reduction in the log of the ratio of ADG between comparison and reference groups. The meta-regression plot is presented in Fig. 4. The predicted impact of increasing difference in FEC between groups on relative daily growth rate performance is presented in Fig. 5. Table 5.

The geographical subset analysis is summarised in Table 3. All continents had meta-analysis risk-ratios less than one and all had negative meta-regression beta coefficients for the common logarithm of difference in FEC between comparison and reference groups (plus 1).

4. Discussion

Whilst some residual heterogeneity remained after combining studies in the meta-analysis, subset analysis indicated that the direction of effect of elevated FEC on reducing ADG was consistent across continents, animal class and production systems. The generalisable impact of strongyle parasitism, as measured by FEC, on daily weight gain of growing cattle is negative and this is independent of continent and production system. Reduced ADG was often present when the

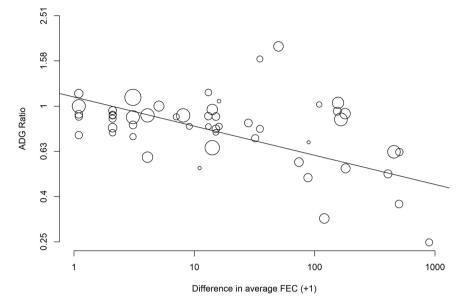


Fig. 4. Meta-regression of log₁₀ faecal egg count difference (+1) between groups on the log of the ratio of average daily weight gain. Point size is linked to the size of the standard error of the ratio, which is determined by the number of animals within each group and the variation in growth rate between individuals within the group.

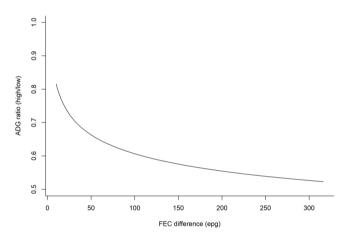


Fig. 5. Predicted average daily weight gain ratio (ADG) by faecal egg count (FEC) difference between comparison and reference group.

Table 5

| Meta-analysis and | l meta-regression | geographical | breakdown. |
|-------------------|-------------------|--------------|------------|
|-------------------|-------------------|--------------|------------|

| Continent | No. studies | No. comp. | Risk ratio | 95% CI | Beta | S. E. | P- value |
|------------------|----------------|--------------|---------------|-------------|--------|----------|-------------|
| Nth America | 15 | 40 | 0.87 | 0.76–0.99 | -0.149 | .07 | 0.03 |
| Europe | 6 | 9 | 0.98 | 0.84 - 1.06 | -0.180 | .08 | 0.03 |
| Sth America | 3 | 8 | 0.67 | 0.52–0.88 | -0.378 | .15 | 0.01 |
| Asia/ oceania | 1 | 1 | 0.76 | 0.68–0.85 | NA | NA | NA |
| Africa | 1 | 1 | 0.76 | 0.36 - 1.61 | NA | NA | NA |
| All | 26 | 59 | 0.88 | 0.82-0.95 | -0.131 | .04 | < 0.01 |

differences in the faecal egg count was only 50 or less between reference and comparison groups. It should be noted that for several observations a higher ADG was recorded in the comparison group (the group with the higher FEC). However, this effect was not observed when the difference in faecal egg count was 100 or more between the comparison and reference groups.

The relationship between the log of the ratio of ADG and the common

logarithm of the difference in FEC was linear. This means the relationship is curvilinear on the original scale. Strongyle parasitism reduces ADG with the lowest growth rates occurring in groups with the highest FEC burden. The reduction in ADG due to increasing strongyle parasite burden identified in this review is consistent with other findings (Beltrán et al., 2020; Charlier et al., 2009; Merlin et al., 2017; Sutherland and Bullen, 2015; Vercruysse and Claerebout, 2001). A meta-analysis of studies across western Europe showed that in untreated first season grazing calves, 53 mobs considered 'clinical' (mean peak of 275 epg) had ADG of 375 g/head/day compared to 32 'subclinical' mobs (mean peak of 100 epg) at 530 g/head/day (Shaw et al., 1997).

Although the egg output of different gastrointestinal nematodes varies (Bailey, 2008), it was beyond the scope of this investigation to quantify the impact on ADG attributable to different species of parasites. A recent meta-analysis identified a negative impact of gastrointestinal nematode infections on weight gain of growing sheep (Mavrot et al., 2015). This study also described a negative correlation between FEC and ADG, regardless of the worm species identified.

This study shows that WEC differences of 100 epg in young grazing cattle are associated with a material depression in ADG (see Fig. 5). The impact of parasite burdens showing worm egg counts below 100epg have previously been demonstrated (George et al., 2020; Stromberg et al., 2012). Practitioners and producers will determine their own thresholds for intervention, depending on the cost of mustering and treatment or other interventions such as supplementing feed, compared with the costs associated with depressed growth rates and pasture larval contamination. Effective control of strongyle parasite burden in growing cattle provides growth benefits. Identifying sustainable and cost-effective strongyle parasite controls remains an important objective for effective management and production of cattle globally. This paper provides evidence to justify investment in interventions for young cattle to minimise exposure to the larvae of, as well as treating excessive burdens of, gastrointestinal nematodes and in doing so improve both productivity and welfare.

Declarations

Ethics approval is not applicable. No animal or human experimentation was undertaken.

Consent to publication is not applicable for the reasons listed above.

All data generated or analysed during this study are available from referenced published articles (meta-analysis).

RWS has no competing interests.

RWS received funding from Zoetis Australia to undertake this work. RWS undertook the literature review, meta-analysis. All authors contributed to study design and writing of the manuscript.

CRediT authorship contribution statement

RWS co-designed the study, undertook the systematic literature review, conducted the analysis and authoring of the manuscript, ASH, MP and SO contributed to study design, assisted with result interpretation and co-authored the paper.

Declaration of Competing Interest

The authors declare the following financial interests/personal relationships which may be considered as potential competing interests: Richard Shephard reports financial support was provided by Zoetis Australia Pty Limited.

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